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# **Identifying missing dictionary entries with frequency-conserving context models**

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In an effort to better understand meaning from natural language texts, we explore methods aimed at organizing lexical objects into contexts. A number of these methods for organization fall into a family defined by word ordering. Unlike demographic or spatial partitions of data, these collocation models are of special importance for their universal applicability. While we are interested here in text and have framed our treatment appropriately, our work is potentially applicable to other areas of research (e.g., speech, genomics, and mobility patterns) where one has ordered categorical data (e.g., sounds, genes, and locations). Our approach focuses on the phrase (whether word or larger) as the primary meaning-bearing lexical unit and object of study. To do so, we employ our previously developed framework for generating word-conserving phrase-frequency data. Upon training our model with the Wiktionary, an extensive, online, collaborative, and open-source dictionary that contains over 100 000 phrasal definitions, we develop highly effective filters for the identification of meaningful, missing phrase entries. With our predictions we then engage the editorial community of the Wiktionary and propose short lists of potential missing entries for definition, developing a breakthrough, lexical extraction technique and expanding our knowledge of the defined English lexicon of phrases.

#### **I. BACKGROUND**

Starting with the work of Shannon [\[1\]](#page-13-0), information theory has grown enormously and has been shown by Jaynes to have deep connections to statistical mechanics [\[2\]](#page-13-0). We focus on a particular aspect of Shannon's work, namely, joint probability distributions between word types (denoted by  $w \in W$ ) and their groupings by appearance orderings, or *contexts* (denoted by  $c \in C$ ). For a word appearing in text, Shannon's model assigned context according to the word's immediate antecedent. In other words, the sequence

#### $\cdots w_{i-1}w_i \cdots$

places this occurrence of the word type of  $w_i$  in the context of *wi*−1*-* (uniquely defined by the word type of *wi*−1), where the star denotes any word. This experiment was novel, and when these transition probabilities were observed, he found a method for the automated production of language that far better resembled true English text than simple adherence to relative word frequencies.

Later, though still early on in the history of modern computational linguistics and natural language processing, theory caught up with Shannon's work. Becker wrote [\[3\]](#page-13-0) the following.

My guess is that phrase-adaption and generative gap-filling are very roughly equally important in language production, as measured in processing time spent on each, or in constituents arising from each. One way of making such an intuitive estimate is simply to listen to what people actually say when they speak. An independent way of gauging the importance of the phrasal lexicon is to determine its size.

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Since then, with the rise of computation and increasing availability of electronic text, there have been numerous extensions of Shannon's context model. These models have generally been information-theoretic applications as well, mainly used to predict word associations [\[4\]](#page-13-0) and to extract multiword expressions (MWEs) [\[5\]](#page-13-0). This latter topic has been one of extreme importance for the computational linguistics community [\[6\]](#page-13-0) and has seen many approaches aside from the information theoretic, including with part-ofspeech taggers [\[7\]](#page-13-0) (where categories, e.g., noun and verb, are used to identify word combinations) and with syntactic parsers [\[8\]](#page-13-0) (where rules of grammar are used to identify word combinations). However, almost all of these methods have the common issue of scalability [\[9\]](#page-13-0), making them difficult to use for the extraction of phrases of more than two words.

Information-theoretic extensions of Shannon's context model have also been used by Piantadosi *et al.* [\[10\]](#page-13-0) to extend the work of Zipf [\[11\]](#page-13-0), using an entropic derivation called the information content (IC)

$$
I(w) = -\sum_{c \in C} P(c|w) \log P(w|c)
$$
 (1)

and measuring its associations with word lengths. Though there have been concerns over some of the conclusions reached in this work [\[12–15\]](#page-13-0), Shannon's model was somewhat generalized and applied to 3-gram, 4-gram, and 5-gram context models to predict word lengths. This model was also used by Garcia *et al.* [\[16\]](#page-13-0) to assess the relationship between sentiment (surveyed emotional response) norms and IC measurements of words. However, their application of the formula

$$
I(w) = -\frac{1}{f(w)} \sum_{i=1}^{f(w)} \log P(w|c_i)
$$
 (2)

to *N*-gram data was wholly incorrect, as this special representation applies only to corpus-level data, i.e., plot line–human readable text, and *not* the frequency-based *N*-grams.

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In addition to the above considerations, there is also the general concern of nonphysicality with imperfect word-frequency conservation, which is exacerbated by the Piantadosi *et al.* extension of Shannon's model. To be precise, for a joint distribution of words and contexts that is *physically* related to the appearance of words on "the page," there should be conservation in the marginal frequencies:

$$
f(w) = \sum_{c \in C} f(w, c),\tag{3}
$$

much like that discussed in [\[4\]](#page-13-0). This property is not upheld using any true, sliding-window *N*-gram data (e.g., [\[17–19\]](#page-13-0)). To see this, we recall that in both [\[16\]](#page-13-0) and [\[10\]](#page-13-0), a word's *N*gram context was defined by its immediate  $N - 1$  antecedents. However, by this formulation we note that the first word of a page appears as last in no 2-gram, the second appears as last in no 3-gram, and so on.

These word-frequency misrepresentations may seem to be of little importance at the text or page level, but since the methods for large-scale *N*-gram parsing have adopted the practice of stopping at sentence and clause boundaries [\[19\]](#page-13-0), wordfrequency misrepresentations (such as those discussed above) have become very significant. In the new format, 40% of the words in a sentence or clause of length 5 have no 3-gram context (the first two). As such, when these context models are applied to modern *N*-gram data, they are incapable of accurately representing the frequencies of words expressed. We also note that despite the advances in processing made in the construction of the current Google *N*-grams corpus [\[19\]](#page-13-0), other issues have been found, namely, regarding the source texts taken [\[20\]](#page-13-0).

The *N*-gram expansion of Shannon's model incorporated more information on relative word placement, but perhaps an ideal scenario would arise when the frequencies of authorintended phrases are exactly known. Here one can conserve word frequencies (as we discuss in Sec.  $II$ ) when a context for an instance of a word is defined by its removal pattern, i.e., the word "cat" appears in the context "∗ in the hat" when the phrase "cat in the hat" is observed. In this way, a word type appears in as many contexts as there are phrase types that contain the word. While we consider the different phrase types as having rigid and different meanings, the words underneath can be looked at as having more flexibility, often in need of disambiguation. This flexibility is quite similar to an aspect of a physical model of lexicon learning [\[21\]](#page-13-0), where a context size control parameter was used to tune the number of plausible but unintended meanings that accompany a single word's true meaning. An enhanced model of lexicon learning that focuses on meanings of phrases could then explain the need for disambiguation when reading by words.

We also note that there exist many other methods for grouping occurrences of lexical units to produce informative context models. Resnik [\[22\]](#page-13-0) showed that class categorizations of words (e.g., verbs and nouns) could be used to produce informative joint probability distributions. Montemurro and Zanette [\[23\]](#page-13-0) used joint distributions of words and arbitrary equal-length parts of texts to entropically quantify the semantic information encoded in written language. Texts tagged with metadata such as genera  $[24]$ , time  $[25]$ , location  $[26]$ , and language [\[27\]](#page-13-0) have rendered straightforward and clear examples of the power in a (word-frequency conserving) joint probability mass function (PMF), shedding light on social phenomena by relating words to classes. Additionally, while their work did not leverage word frequencies or the joint PMFs possible, Benedetto *et al.* [\[28\]](#page-13-0) used metadata of texts to train language and authorship detection algorithms and further to construct accurate phylogeneticlike trees through application of compression distances. Though metadata approaches to context are informative, with their power there is simultaneously a loss of applicability (metadata is frequently not present) as well as a loss of biocommunicative relevance (humans are capable of inferring social information from text in isolation).

#### **II. FREQUENCY-CONSERVING CONTEXT MODELS**

In previous work [\[29\]](#page-13-0) we developed a scalable and general framework for generating frequency data for *N*-grams, called random text partitioning. Since a phrase-frequency distribution *S* is balanced with regard to its underlying word-frequency distribution *W*,

$$
\sum_{w \in W} f(w) = \sum_{s \in S} \ell(s) f(s) \tag{4}
$$

(where  $\ell$  denotes the phrase-length norm, which returns the length of a phrase in numbers of words), it is easy to produce a symmetric generalization of Shannon's model that integrates all phrase or *N*-gram lengths and all word placement or removal points. To do so, we define *W* and *S* to be the sets of words and (text-partitioned) phrases from a text, respectively, and let *C* be the collection of all single-word-removal patterns from the phrases of *S*. A joint frequency  $f(w,c)$  is then defined by the partition frequency of the phrase that is formed when *c* and *w* are composed. In particular, if *w* composed with *c* renders *s*, we then set  $f(w, c) = f(s)$ , which produces a context model on the words whose marginal frequencies preserve their original frequencies from the page. In particular, we refer to this, or such a model for phrases, as an external context model since the relations are produced by structure external to the semantic unit.

It is good to see the external word-context generalization emerge, but our interest actually lies in the development of a context model for the phrases themselves. To do so, we define the internal contexts of a phrase by the patterns generated through the removal of subphrases. To be precise, for a phrase *s* and a subphrase  $s_{i\cdots j}$  ranging over words *i* through *j* we define the context

$$
c_{i\cdots j} = w_1 \cdots w_{i-1} \star \cdots \star w_{j+1} \cdots w_{\ell(s)} \tag{5}
$$

to be the collection of same-length phrases whose analogous word removal (*i* through *j*) renders the same pattern (when word types are considered). We present the contexts of generalized phrases of lengths 1–4 in Table [I,](#page-3-0) as described above. Looking at the table, it becomes clear that these contexts are actually a mathematical formalization of the generative gap filling proposed in [\[3\]](#page-13-0), which was semiformalized by the phrasal templates discussed at length by Smadja in [\[5\]](#page-13-0). Between our formulation and that of Smadja, the main difference of definition lies in our restriction to contiguous word sequence (i.e., subphrase) removals, as is necessitated by the mechanics of the secondary partition process, which defines the context lists.

<span id="page-3-0"></span>TABLE I. Expansion of context lists for longer and longer phrases. We define the internal contexts of phrases by the removal of individual subphrases. These contexts are represented as phrases with words replaced by stars. Any phrases whose word types match after analogous subphrase removals share the matching context. Here the columns are labeled  $1-4$  by subphrase length.

Phrase	$\ell(s_{i\cdots j})=1$	$\ell(s_{i\cdots j})=2$	$\ell(s_{i\cdots j})=3$	$\ell(s_{i\cdots j})=4$	$\cdots$
$w_1$	$\star$				$\ldots$ .
$w_1$ $w_2$	$\star w_2$	$\star \star$			$\cdots$
	$w_1 \star$				$\cdots$
$w_1 w_2 w_3$	$\star w_2 w_3$	$\star \star w_3$	* * *		$\cdots$
	$w_1 \star w_3$	$w_1 \star \star$			$\cdots$
	$w_1 w_2 \star$				$\ldots$
$w_1 w_2 w_3 w_4$	$\star w_2 w_3 w_4$	$\star \star w_3 w_4$	$\star \star \star w_4$	$\star \star \star \star$	$\ldots$
	$w_1 \star w_3 w_4$	$w_1 \star \star w_4$	$w_1 \star \star \star$		$\ldots$ .
	$w_1 w_2 \star w_4$	$w_1 w_2 \star \star$			$\ldots$
	$w_1 w_2 w_3 \star$				$\cdots$
$\vdots$	$\bullet$ $\bullet$ $\bullet$	٠ $\bullet$ $\bullet$	٠ $\cdot$	$\bullet$	$\mathcal{L}_{\mathcal{A}}$

The weighting of the contexts for a phrase is accomplished simultaneously with their definition through a secondary partition process describing the inner contextual modes of interpretation for the phrase. The process is as follows. In an effort to relate an observed phrase to other known phrases, the observer selectively ignores a subphrase of the original phrase. To retain generality, we do this by considering the *random* partitions of the original phrase and then assuming that a subphrase is ignored from a partition with probability proportional to its length, to preserve word (and hence phrase) frequencies. The conditional probabilities of inner context are then

 $P(c_{i\cdots j}|s) = P(\text{ignore } s_{i\cdots j} \text{ given a partition of } s)$ 

= 
$$
P
$$
(ignore  $s_{i\cdots j}$  given  $s_{i\cdots j}$  is partitioned from  $s$ )

$$
\times P(s_{i\cdots j} \text{ is partitioned from } s). \tag{6}
$$

Utilizing the partition probability and our assumption, we note from our work in [\[29\]](#page-13-0) that

$$
\ell(s) = \sum_{1 \leq i < j \leq \ell(s)} \ell(s_{i\cdots j}) P_q(s_{i\cdots j} \mid s),\tag{7}
$$

which ensures through defining

$$
P(c_{i\cdots j}|s) = \frac{\ell(s_{i\cdots j})}{\ell(s)} P_q(s_{i\cdots j}|s),
$$
 (8)

the production of a valid, phrase-frequency preserving context model

$$
\sum_{c \in C} f(c,s) = \sum_{i < j \le \ell(s)} P(c_{i\cdots j}|s) f(s)
$$
\n
$$
= f(s) \sum_{1 \le i < j \le \ell(s)} \frac{\ell(s_{i\cdots j})}{\ell(s)} P_q(s_{i\cdots j}|s) = f(s), \quad (9)
$$

which preserves the underlying frequency distribution of phrases. Note here that beyond this point in the paper we will used the normalized form

$$
P(c,s) = \frac{f(c,s)}{\sum_{s \in S} \sum_{c \in C} f(c,s)}\tag{10}
$$

for convenience in the derivation of expectations in the next section.

# **III. LIKELIHOOD OF DICTIONARY DEFINITION**

In this section we exhibit the power of the internal context model through a lexicographic application, deriving a measure of meaning and definition for phrases with empirical phrase-definition data taken from a collaborative open-access dictionary [\[30\]](#page-13-0) (see Sec. [V](#page-4-0) for more information on our data and the Wiktionary). With the rankings that this measure derives, we will go on to propose phrases for definition with the editorial community of the Wiktionary in an ongoing live experiment, discussed in Sec. [IV.](#page-4-0)

To begin, we define the dictionary indicator *D* to be a binary norm on phrases, taking value 1 when a phrase appears in the dictionary (i.e., has definition) and taking value 0 when a phrase is unreferenced. The dictionary indicator tells us when a phrase has reference in the dictionary and in principle can be replaced with other indicator norms, for other purposes. Moving forward, we take note of an intuitive description of the distribution average

$$
\overline{D}(S) = \sum_{t \in S} D(t)P(t)
$$
  
=  $P(\text{randomly drawing a defined phrase from } S)$ 

and go on to derive an alternative expansion through application of the context model

$$
\overline{D}(S) = \sum_{t \in S} D(t)P(t)
$$
\n
$$
= \sum_{t \in S} D(t)P(t) \sum_{c \in C} P(c|t) \sum_{s \in S} P(s|c)
$$
\n
$$
= \sum_{c \in C} P(c) \sum_{t \in S} D(t)P(t|c) \sum_{s \in S} P(s|c)
$$
\n
$$
= \sum_{c \in C} P(c) \sum_{s \in S} P(s|c) \sum_{t \in S} D(t)P(t|c)
$$
\n
$$
= \sum_{s \in S} P(s) \sum_{c \in C} P(c|s) \sum_{t \in S} D(t)P(t|c)
$$
\n
$$
= \sum_{s \in S} P(s) \sum_{c \in C} P(c|s) \overline{D}(c|S). \tag{11}
$$

<span id="page-4-0"></span>In the last line we then interpret

$$
\overline{D}(C|s) = \sum_{c \in C} P(c|s)\overline{D}(c|S)
$$
 (12)

to be the likelihood [analogous to the IC equation presented here as Eq. [\(1\)](#page-1-0)] that a phrase, which is randomly drawn from a context of *s*, to have definition in the dictionary. To be precise, we say  $D(C|s)$  is the likelihood of dictionary definition of the context model *C*, given the phrase *s*, or, when only one *c* ∈ *C* is considered, we say  $D(c|S) = \sum_{t \in S} D(t)P(t|c)$  is the likelihood of dictionary definition of the context *c*, given *S*. Numerically, we note that the distribution-level values *D*(*C*|*s*) extend the dictionary over all *S*, smoothing out the binary data to the full lexicon (uniquely for phrases of more than one word, which have no interesting space-defined internal structure) through the relations of the model. In other words, though  $\overline{D}(C|s) \neq 0$  may now only indicate the *possibility* of a phrase having definition, it is still a strong indicator and most importantly may be applied to never-before-seen expressions. We illustrate the extension of the dictionary through  $\overline{D}$  in Fig. 1, where it becomes clear that the topological structure of the associated network of contexts is crystalline, unlike the small-world phenomenon observed for the words of a thesaurus in [\[31\]](#page-13-0). However, this is not surprising, given that the latter is a conceptual network defined by common meanings, as opposed to a rigid, physical property, such as word order.

#### **IV. PREDICTING MISSING DICTIONARY ENTRIES**

Starting with the work of Sinclair *et al.* [\[32\]](#page-13-0) (though the idea was proposed more than ten years earlier by Becker in [\[3\]](#page-13-0)), lexicographers have been building dictionaries based on language as it is spoken and written, including idiomatic, slang-filled,



FIG. 1. (Color online) Example showing the sharing of contexts by similar phrases. Suppose that our text consists of the two phrases, "in the contrary" and "on the contrary," that each occurs once, and that the latter has definition  $(D = 1)$  while the former does not. In this event, we see that the three shared contexts " $\star \star \star$ ", " $\star \star$  contrary," and " $\star$  the contrary" present elevated likelihood *D* values, indicating that the phrase "in the contrary" may have meaning and be worthy of definition.

and grammatical expressions [\[33–36\]](#page-13-0). These dictionaries have proven highly effective for nonprimary language learners, who may not be privy to cultural metaphors. In this spirit, we utilize the context model derived above to discover phrases that are undefined, but which may be in need of definition for their similarity to other, defined phrases. We do this in a corpus-based way, using the definition likelihood  $\overline{D}(C|s)$ as a secondary filter to frequency. The process is in general quite straightforward and first requires a ranking of phrases by frequency of occurrence  $f(s)$ . Upon taking the first  $s_1, \ldots, s_N$ frequency-ranked phrases  $(N = 100000,$  for our experiments), we reorder the list according to the values  $D(C|s)$ (descending). The top of such a double-sorted list then includes phrases that are both frequent and similar to defined phrases.

With our double-sorted lists we then record those phrases having no definition or dictionary reference, but which are at the top. These phrases are quite often meaningful (as we have found experimentally; see below) despite their lack of definition and as such we propose this method for the automated generation of short lists for editorial investigation of definition.

#### **V. MATERIALS AND METHODS**

For its breadth, open-source nature, and large editorial community, we utilize dictionary data from the Wiktionary [\[30\]](#page-13-0) (a Wiki-based open content dictionary) to build the dictionaryindicator norm, setting  $D(s) = 1$  if a phrase *s* has reference or redirect. We apply our filter for missing entry detection to several large corpora from a wide scope of content. These corpora are 20 years of*New York Times* articles (*NYT*, 1987–2007) [\[37\]](#page-13-0), approximately 4% of a year's tweets (Twitter, 2009) [\[38\]](#page-13-0), music lyrics from thousands of songs and authors (lyrics, 1960–2007) [\[24\]](#page-13-0), complete Wikipedia articles (Wikipedia, 2010) [\[39\]](#page-13-0), and Project Gutenberg eBooks collection (eBooks, 2012) [\[40\]](#page-13-0) of more than 30 000 public-domain texts. We note that these are all unsorted texts and that Twitter, eBooks, Lyrics, and to an extent Wikipedia are mixtures of many languages (though the majority is English). We only attempt missing entry prediction for phrase lengths (2–5), for their inclusion in other major collocation corpora [\[19\]](#page-13-0), and their having the most data in the dictionary. We also note that all text processed is taken lowercase.

To understand our results, we perform a tenfold cross-validation on the frequency and likelihood filters. This is executed by randomly splitting the Wiktionary's list of defined phrases into ten equal-length pieces and then performing ten parallel experiments. In each of these experiments we determine the likelihood values  $\overline{D}(C|s)$  by a distinct  $\frac{9}{10}$  of the data. We then order the union set of the  $\frac{1}{10}$ -withheld and the Wiktionary-undefined phrases by their likelihood (and frequency) values descending and accept some top segment of the list, or short list, coding them as positive by the experiment. For such a short list, we then record the true positive rates, i.e., portion of all  $\frac{1}{10}$ -withheld truly defined phrases we coded positive, the false positive rates, i.e., portion of all truly undefined phrases we coded positive, and the number of entries discovered. Upon performing these experiments, the average of the ten trials is taken for each of the three parameters, for a number of short list lengths (scanning 1000 log-spaced lengths), and plotted as a receiver operating characteristic (ROC) curve (see Figs. 2[–6\)](#page-8-0). We also note that each is also presented with its area under curve (AUC), which measures the accuracy of the expanding-list classifier as a whole.

#### **VI. RESULTS AND DISCUSSION**

Before observing output from our model we take the time to perform a cross-validation (tenfold) and compare our context filter to a sort by frequency alone (see Fig. 2 below and Figs. [3](#page-7-0)[–6](#page-8-0) in Appendix [A\)](#page-7-0). From this we have found that our likelihood filter renders missing entries much more efficiently than by frequency (see Table II and Figs.  $2-6$ ), already discovering missing entries from short lists of as little as 20 (see the insets of Figs.  $2-6$  as well as Tables II[–VII\)](#page-12-0). As such we adhere to this standard and only publish short lists of 20 predictions per corpus per phrase lengths 2–5. In parallel, we also present phrase frequency-generated short lists for comparison.

In addition to listing them in Appendix  $\bf{B}$ , we have presented the results of our experiment from across the five large disparate corpora on the Wiktionary in a pilot program, where



FIG. 2. (Color online) With data taken from the Twitter corpus, we present (tenfold) cross-validation results for the filtration procedures. For each of the lengths 2, 3, 4, and 5, we show the ROC curves, comparing true and false positive rates for both the likelihood filters (black) and the frequency filters (gray). There we see increased performance in the likelihood classifiers (except possibly for length 5), which is reflected in the AUCs (where an AUC of 1 indicates a perfect classifier). In the insets we also monitor the average number of missing entries discovered as a function of the number of entries proposed, for each length. There the horizontal dotted lines indicate the average numbers of missing entries discovered for both the likelihood filters (black) and the frequency filters (gray) when short lists of 20 phrases were taken (red dotted vertical lines). From this we see an indication that even the 5-gram likelihood filter is effective at detecting missing entries in short lists, while the frequency filter is not.

TABLE II. Summarizing our results from the cross-validation procedure (top), we present the mean numbers of missing entries discovered when 20 guesses were made for *N*-grams or phrases of lengths 2, 3, 4, and 5 each. For each of the five large corpora (see Sec. [V\)](#page-4-0) we make predictions according our likelihood filter and according to frequency (in parentheses) as a baseline. When considering the 2-grams (for which the most definition information exists), short lists of 20 rendered up to 25% correct predictions on average by the definition likelihood, as opposed to the frequency ranking, by which no more than 2*.*5% could be expected. We also summarize the results to date from the live experiment (bottom) (updated 19 February 2015) and present the numbers of missing entries correctly discovered on the Wiktionary (i.e., reference added since 1 July 2014, when the dictionary's data was accessed) by the 20-phrase short lists produced in our experiments for both the likelihood and frequency (in parentheses) filters. Here we see that all of the corpora analyzed were generative of phrases, with Twitter far and away being the most productive and the reference corpus Wikipedia the least so.



we are tracking the success of the filters [\[41\]](#page-13-0). Looking at the lexical tables, where defined phrases are highlighted bold, we can see that many of the predictions by the likelihood filter (especially those obtained from the Twitter corpus) have already been defined in the Wiktionary following our recommendation [\[30\]](#page-13-0). We also summarize these results from the live experiment in Table II.

Looking at the lexical tables more closely, we note that all corpora present highly idiomatic expressions under the likelihood filter, many of which are variants of existing idiomatic phrases that will likely be granted inclusion into the dictionary through redirects or alternative-form listings. To name a few, the Twitter (Table [III\)](#page-6-0), *NYT* (Table [IV\)](#page-9-0), and lyrics (Table [V\)](#page-10-0) corpora consistently predict large families derived from phrases such as "at the same time" and "you know what i mean," while the eBooks and Wikipedia corpora predict families derived from phrases such as "on the other hand" and "at the same time." In general, we see no such structure or predictive power emerge from the frequency filter.

We also observe that from those corpora that are less pure of English context (namely, the eBooks, lyrics, and Twitter corpora), extra-English expressions have crept in. This highlights an important feature of the likelihood filter: It does not intrinsically rely on the syntax or grammar of the language

<span id="page-6-0"></span>TABLE III. With data taken from the Twitter corpus, we present the top 20 unreferenced phrases considered for definition (in the live experiment) from each of the 2-, 3-, 4-, and 5-gram likelihood filters (top) and frequency filters (bottom). From this corpus we note the juxtaposition of highly idiomatic expressions by the likelihood filter (such as "holy #!@&"), with the domination of the frequency filters by semiautomated content. The phrase "holy #!@&" is an example of the model's success with this corpus, as it achieved definition several months after the Wiktionary's data were accessed.



The symbols used in Tables III and [V](#page-10-0) represent the words shit =  $@**\$ , ass =  $\frac{1}{6}$ , fuck =  $\&$  \*#!, and hell = #! $@\&$ .

to which it is applied, beyond the extent to which syntax and grammar effect the shapes of collocations. For example, the eBooks predict (see Table [VII\)](#page-12-0) the undefined French phrase "tu ne sais pas" or "you do not know," which is a syntactic variant of the English-Wiktionary defined French, "je ne sais pas," meaning "i do not know." Seeing this, we note that it would be straightforward to construct a likelihood filter with a language indicator norm to create an alternative framework for language identification.

There are also a fair number of phrases predicted by the likelihood filter that in fact are spelling errors, typographical errors, and grammatical errors. In terms of the context model, these erroneous forms are quite close to those defined in the dictionary and so rise in the short lists generated from the lesswell-edited corpora, e.g., "actions speak louder *then* words" in the Twitter corpus. This then seems to indicate the potential for the likelihood filter to be integrated into autocorrect algorithms and further points to the possibility of constructing

<span id="page-7-0"></span>syntactic indicator norms of phrases, making estimations of tenses and parts of speech (whose data are also available from the Wiktionary [\[30\]](#page-13-0)) possible through application of the model in precisely the same manner presented in Sec. [III.](#page-3-0) Regardless of the future applications, we have developed and presented a powerful and scalable MWE extraction technique.

#### **APPENDIX A: CROSS-VALIDATION RESULTS FOR MISSING ENTRY DETECTION**

In this Appendix we provide cross-validation results for missing entry detection.

#### **1. The New York Times**



FIG. 3. (Color online) With data taken from the *NYT* corpus, we present (tenfold) cross-validation results for the filtration procedures. For each of the lengths 2, 3, 4, and 5, we show the ROC curves, comparing true and false positive rates for both the likelihood filters (black) and the frequency filters (gray). There we see increased performance in the likelihood classifiers (except possibly for length 5), which is reflected in the AUCs (where an AUC of 1 indicates a perfect classifier). In the insets we also monitor the average number of missing entries discovered as a function of the number of entries proposed, for each length. There the horizontal dotted lines indicate the average numbers of missing entries discovered for both the likelihood filters (black) and the frequency filters (gray) when short lists of 20 phrases were taken (red dotted vertical lines). From this we see an indication that even the 5-gram likelihood filter is effective at detecting missing entries in short lists, while the frequency filter is not.





FIG. 4. (Color online) With data taken from the lyrics corpus, we present (tenfold) cross-validation results for the filtration procedures. For each of the lengths 2, 3, 4, and 5, we show the ROC curves, comparing true and false positive rates for both the likelihood filters (black) and the frequency filters (gray). There we see increased performance in the likelihood classifiers, which is reflected in the AUCs (where an AUC of 1 indicates a perfect classifier). In the insets we also monitor the average number of missing entries discovered as a function of the number of entries proposed, for each length. There the horizontal dotted lines indicate the average numbers of missing entries discovered for both the likelihood filters (black) and the frequency filters (gray), when short lists of 20 phrases were taken (red dotted vertical lines). Here we can see that it may have been advantageous to construct slightly longer 3- and 4-gram lists.

#### **3. English Wikipedia**

# **4. Project Gutenberg eBooks**

<span id="page-8-0"></span>

FIG. 5. (Color online) With data taken from the Wikipedia corpus, we present (tenfold) cross-validation results for the filtration procedures. For each of the lengths 2, 3, 4, and 5, we show the ROC curves, comparing true and false positive rates for both the likelihood filters (black) and the frequency filters (gray). There we see increased performance in the likelihood classifiers, which is reflected in the AUCs (where an AUC of 1 indicates a perfect classifier). In the insets we also monitor the average number of missing entries discovered as a function of the number of entries proposed, for each length. There the horizontal dotted lines indicate the average numbers of missing entries discovered for both the likelihood filters (black) and the frequency filters (gray) when short lists of 20 phrases were taken (red dotted vertical lines). Here we can see that it may have been advantageous to construct slightly longer 3- and 4-gram lists.



FIG. 6. (Color online) With data taken from the eBooks corpus, we present (tenfold) cross-validation results for the filtration procedures. For each of the lengths 2, 3, 4, and 5, we show the ROC curves, comparing true and false positive rates for both the likelihood filters (black) and the frequency filters (gray). There we see increased performance in the likelihood classifiers, which is reflected in the AUCs (where an AUC of 1 indicates a perfect classifier). In the insets we also monitor the average number of missing entries discovered as a function of the number of entries proposed, for each length. There the horizontal dotted lines indicate the average numbers of missing entries discovered for both the likelihood filters (black) and the frequency filters (gray) when short lists of 20 phrases were taken (red dotted vertical lines). Here we can see that the power of the 4-gram model does not show itself until longer lists are considered.

## **APPENDIX B: TABLES OF POTENTIAL MISSING ENTRIES**

<span id="page-9-0"></span>In this Appendix we provide lexical tables of potential missing entries.

#### **1. The New York Times**

TABLE IV. With data taken from the *NYT* corpus, we present the top 20 unreferenced phrases considered for definition (in the live experiment) from each of the 2-, 3-, 4-, and 5-gram likelihood filters (top) and frequency filters (bottom). From this corpus we note the juxtaposition of highly idiomatic expressions by the likelihood filter (such as "united front"), with the domination of the frequency filters by structural elements of rigid content (e.g., the obituaries). The phrase "united front" is an example of the model's success with this corpus, as its coverage in a Wikipedia article began in 2006, describing the general Marxist tactic extensively. We also note that we have abbreviated "national oceanographic and atmospheric administration" (column 5, row 2), for brevity.



### **2. Music lyrics**

<span id="page-10-0"></span>TABLE V. With data taken from the lyrics corpus, we present the top 20 unreferenced phrases considered for definition (in the live experiment) from each of the 2-, 3-, 4-, and 5-gram likelihood filters (top) and frequency filters (bottom). From this corpus we note the juxtaposition of highly idiomatic expressions by the likelihood filter (such as "iced up"), with the domination of the frequency filters by various onomatopoeias. The phrase "iced up" is an example of the model's success with this corpus, having had definition in the Urban Dictionary since 2003, indicating that one is "covered in diamonds." Further, though this phrase does have a variant that is defined in the Wiktionary (as early as 2011)—"iced out"—we note that the reference is also made in the Urban Dictionary (as early as 2004), where the phrase has distinguished meaning for one that is so bedecked—ostentatiously.



The symbols used in Tables [III](#page-6-0) and V represent the words shit =  $@^{*}\hat{ }$ , ass = !%&, fuck = &\*#!, and hell = #!@&.

#### **3. English Wikipedia**

TABLE VI. With data taken from the Wikipedia corpus, we present the top 20 unreferenced phrases considered for definition (in the live experiment) from each of the 2-, 3-, 4-, and 5-gram likelihood filters (top) and frequency filters (bottom). From this corpus we note the juxtaposition of highly idiomatic expressions by the likelihood filter (such as "same-sex couples"), with the domination of the frequency filters by highly descriptive structural text from the presentations of demographic and numeric data. The phrase "same-sex couples" is an example of the model's success with this corpus and appears largely because of the existence of the distinct phrases "same-sex marriage" and "married couples" with definitions in the Wiktionary.



#### **4. Project Gutenberg eBooks**

<span id="page-12-0"></span>TABLE VII. With data taken from the eBooks corpus, we present the top 20 unreferenced phrases considered for definition (in the live experiment) from each of the 2-, 3-, 4-, and 5-gram likelihood filters (top) and frequency filters (bottom). From this corpus we note the juxtaposition of many highly idiomatic expressions by the likelihood filter, with the domination of the frequency filters by highly structural text. Here, since the texts are all within the public domain, we see that this much less modern corpus is without the innovation present in the other, but that the likelihood filter does still extract many unreferenced variants of Wiktionary-defined idiomatic forms.



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